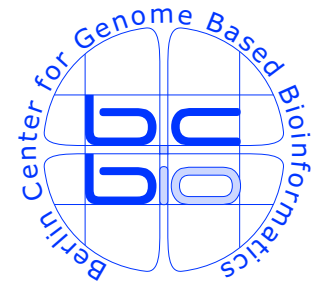
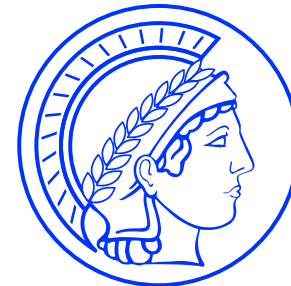


# Computational Inference of Cellular Pathways

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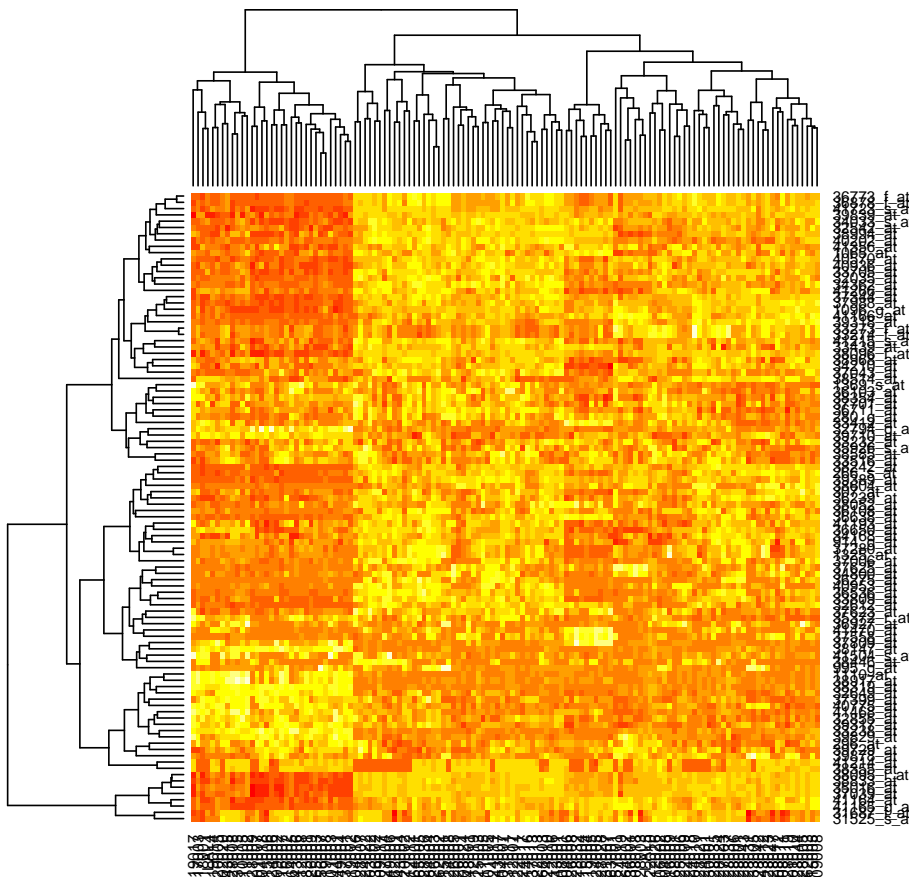


**Practical Microarray Analysis 2005**

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# Coexpression



Coexpression hints to

- **coregulation**
- **gene function**

If genes show the same expression profiles they follow the same regulatory regimes [4, 10].



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# Co-expression graphs

An expression profile is a random vector  $\mathbf{X} = (X_1, \dots, X_p)$ .

**Correlation graph:** Depict genes as vertices of a graph and draw an edge  $(i, j)$  iff the correlation coefficient  $\rho_{ij} \neq 0$ .

**Advantage:** This representation of the marginal dependence structure is **easy to interpret** and can be **accurately estimated** even if  $p \gg N$ .

**Application:** Stuart *et. al* [11] build a graph from coexpression across multiple organisms.■

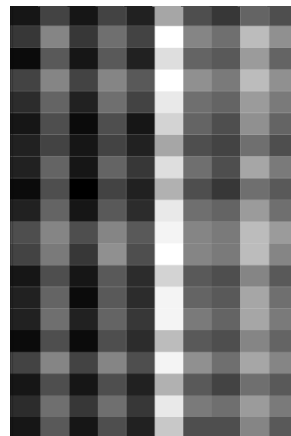


# Differential Co-expression

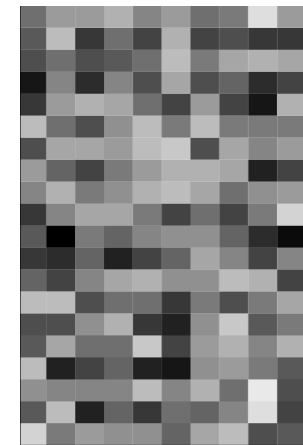
Kostka and Spang [6] find sets of genes, which are correlated in one environment and lose correlation in the second environment.

**Interpretation:** loss (or gain) of regulatory mechanism.

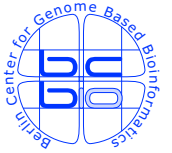
Genes are controlled by common regulatory mechanism:



Genes in chaos!  
No regulatory mechanism to organize expression:



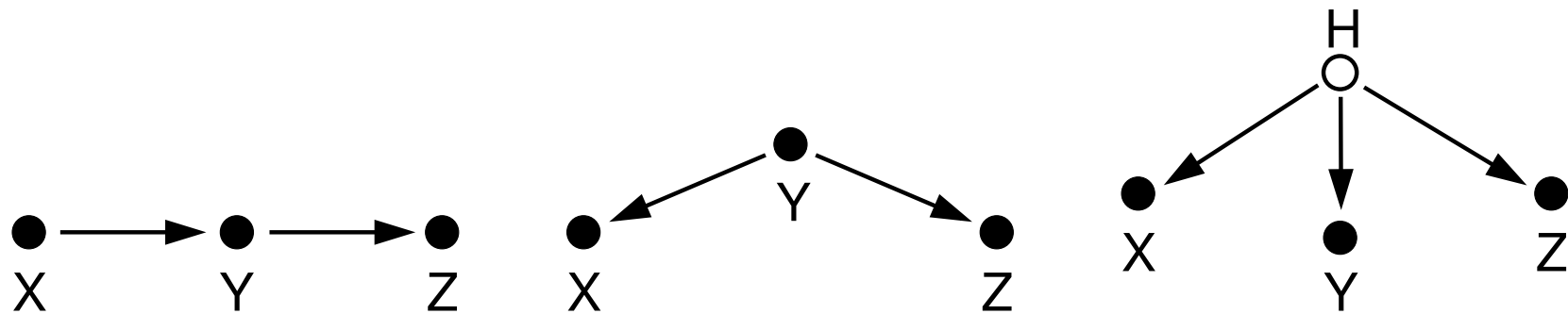
R-package `dcoex` in preparation.



# Problems of co-expression based approaches

**We cannot distinguish direct from indirect dependencies!**

Three reasons, why  $X$ ,  $Y$ , and  $Z$  are highly correlated:



**As a cure:** search for correlations which cannot be explained by other variables.■



# Part I.

# Conditional independence models



---

# Conditional independence

Be  $X, Y, Z$  random variables with joint distribution  $P$ .

**$X$  is conditionally independent of  $Y$  given  $Z$**

$$X \perp\!\!\!\perp Y \mid Z \Leftrightarrow$$

$$P(X = x, Y = y \mid Z = z) = P(X = x \mid Z = z) \cdot P(Y = y \mid Z = z) \blacksquare$$

**If I already know  $Z$ ,  
then  $Y$  offers me no new information  
to understand  $X$ .**

# Conditional independence in Gaussian models

Assume that genes are multivariate normal distributed with covariance matrix  $\Sigma$ . We call  $K = \Sigma^{-1}$  the *concentration matrix* of the distribution.

Then it holds for two genes  $i$  and  $j$ :

$$X_i \perp\!\!\!\perp X_j \mid \mathbf{X}_{\text{rest}} \Leftrightarrow k_{ij} = 0$$

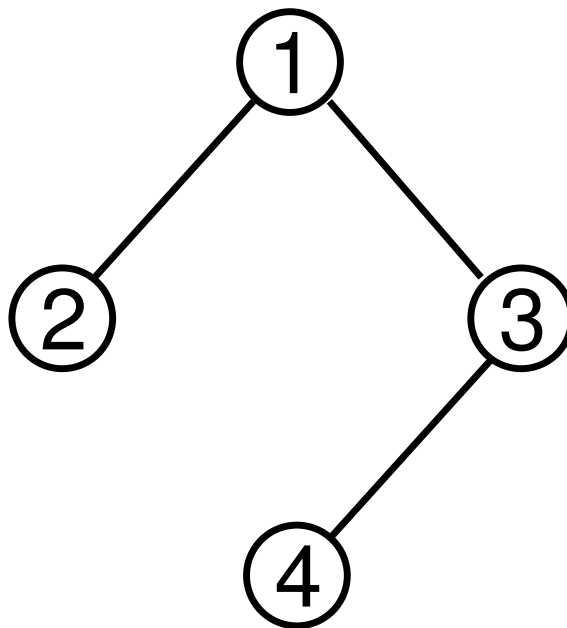
Coexpression networks model via the correlation matrix  $\Sigma$ ,  
Gaussian Graphical Models (GGMs) use the inverse  $K = \Sigma^{-1}$ . ■





# Gaussian Graphical Model

Missing edges indicate independencies:



$$X_i \perp\!\!\!\perp X_j \mid \mathbf{X}_{\text{rest}}$$

$$X_1 \perp\!\!\!\perp X_4 \mid \{X_2, X_3\}$$

$$X_2 \perp\!\!\!\perp X_3 \mid \{X_1, X_4\}$$

$$X_2 \perp\!\!\!\perp X_4 \mid \{X_1, X_3\}$$



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## What if $p \gg N$ ?

Full conditional relationships can only be accurately estimated if the number of samples  $N$  is relatively large compared to the number of variables  $p$ .

Thus, if  $p \gg N$ , you can . . .

- 1. either** improve your estimators of partial correlations (e.g. Schäfer and Strimmer [9] use the **Moore-Penrose pseudoinverse** and **bootstrap aggregation** (bagging) to stabilize the estimator.)
- 2. or** resort to a simpler model.



---

# Sparse graphical modeling

**Idea:** Do not condition on the complete rest as in GGMs. Instead explore dependency of two variables given a single third one.

Draw an edge between genes  $i$  and  $j$  if they are correlated and no third variable can explain the correlation:

$$X_i \not\perp\!\!\!\perp X_j \mid X_k \quad \text{for all } k \in \text{rest.}$$

## Implementations

sparse GGMs [2, 7, 12, 13]; mutual information: ARACNE [1]





## Where are we?

We have seen methods to build graphs from

1. marginal dependencies

$$X_i \perp\!\!\!\perp X_j \mid \emptyset$$

2. full conditional dependence

$$X_i \perp\!\!\!\perp X_j \mid X_{\text{rest}}$$

3. first order dependencies

$$X_i \perp\!\!\!\perp X_j \mid X_k \quad \forall k \in \text{rest}$$

4. This leads use to include **all higher order dependencies**

$$X_i \perp\!\!\!\perp X_j \mid \mathbf{X}_S \quad \text{for all } S \subseteq \text{rest}$$



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# Bayesian network

A Bayesian Network for a random vector  $\mathbf{X}$  consists of

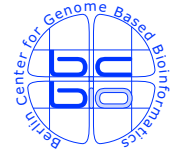
## 1. a network structure

- directed acyclic graph (DAG) on vertex set  $V$ ,
- node  $v$  corresponds to variable  $X_v$ ,

## 2. a set of local probability distributions

- conditional distribution of a gene given its parents.

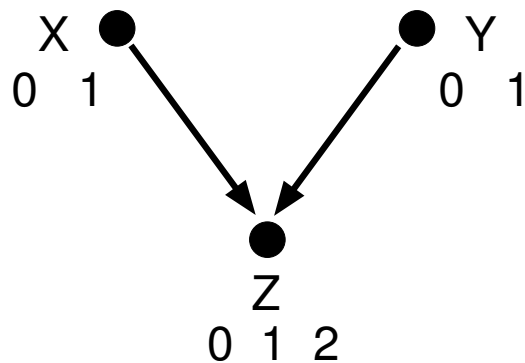
$$p(\mathbf{x}) = \prod_{v \in V} p(x_v \mid \mathbf{x}_{pa(v)}, \theta_v)$$



# Children depend on parents



The DAG defines families.  
Relationships are further characterized by local probability distributions:



$$p(x) = (0.6 \quad 0.4)$$

$$p(y) = (0.2 \quad 0.8)$$

$$p(z|x, y) = \begin{cases} (0.8 & 0.1 & 0.1) & \text{if } (X, Y) = (0, 0) \\ (0.1 & 0.8 & 0.1) & \text{if } (X, Y) = (0, 1) \\ (0.1 & 0.8 & 0.1) & \text{if } (X, Y) = (1, 0) \\ (0.1 & 0.1 & 0.8) & \text{if } (X, Y) = (1, 1) \end{cases}$$





---

## A caveat [5]

If the expression of gene  $A$  is regulated by proteins  $B$  and  $C$ , then  $A$ 's expression level is a function of the joint activity levels of  $B$  and  $C$ .

We treat the expression of  $A$  as a stochastic function of its regulators.■

**Problem 1:** In most current biological data sets, however, we do not have access to measurements of protein activity levels.■

**Resort:** Expression levels of genes as a proxy for the activity level of the proteins they encode.■

**Problem 2:** There are numerous examples where an activation or silencing of a regulator is carried out by posttranscriptional protein modifications.■





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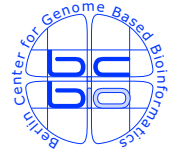
## A first summary

1. Conditional independence is the central concept of statistical network models;
2. Graphical models ask: “Can the correlation between two genes be attributed to other genes?”
3. Increasing order of resolution:  
Clustering, Graphical Gaussian models, Bayesian networks;
4. Models don't capture signaling on protein level.



## Part II.

# Learning from interventions



# Motivation

## Response to microbial challenge

(Boutros *et al.*, Dev Cell, 2002)

Columns: silenced genes.

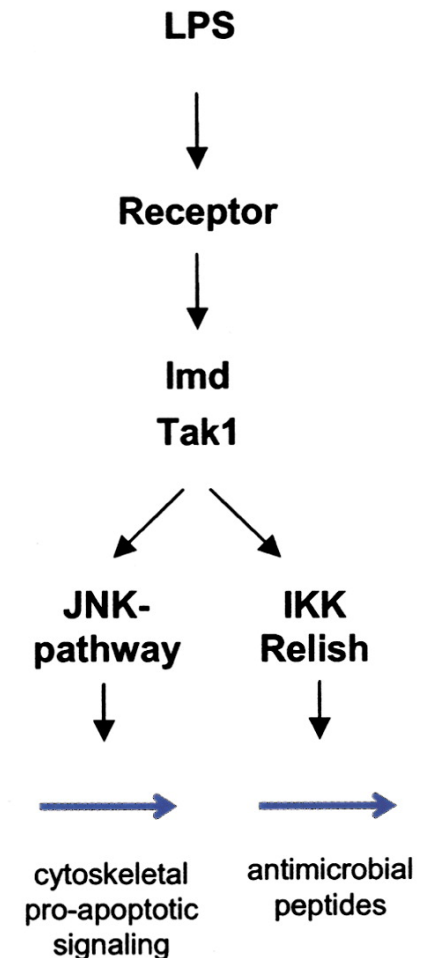
Rows: effects on other genes.

### Results:

1. Silencing **tak1** reduces expression of all LPS-inducible transcripts.
2. Silencing **rel** (**key**) or **mkk4/hep** reduces expression of separate sets of induced transcripts.

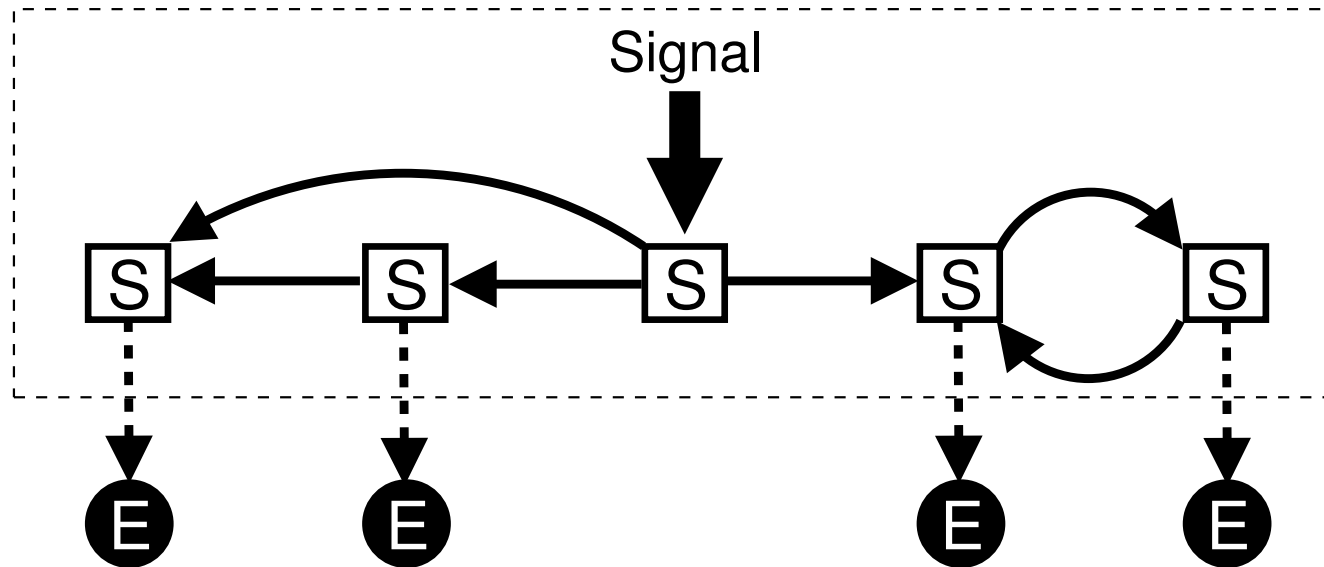


Figures from (Boutros *et al.*, 2002)





## The model [8]



**S-genes** (for “signaling” or “silenced”): candidate pathway genes.

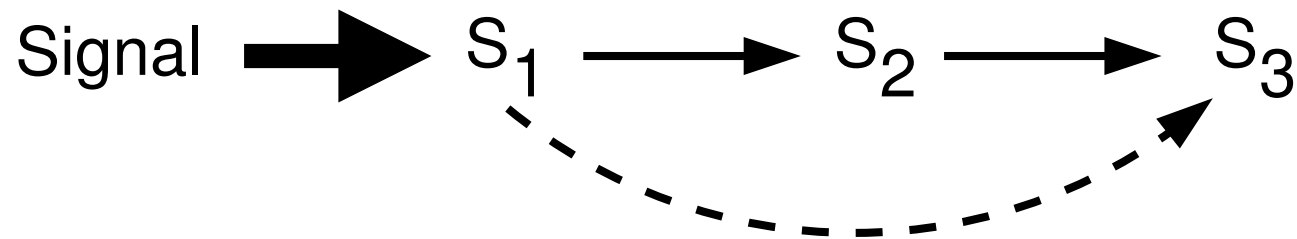
**E-genes** (for “effects”): reporters for S-gene activity.



# Silencing schemes

A pathway topology allows prediction of intervention effects.

We summarize predictions in a **silencing scheme**  $\Phi$ : also a directed graph on S-genes, but transitively closed.



Framework flexible to include **epistatic effects** by local logics.



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# Experiments and Data

Do microarrays for:

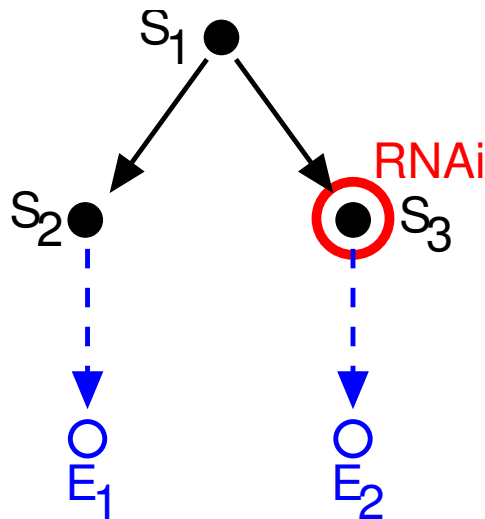
- 1. Negative controls** no signal, no interventions
- 2. Positive controls** pathway activated by signal, no interventions
- 3. Interventions** while signal is on!

**Data:** binary matrix  $D = (e_{ik})$ ,

where  $e_{ik} = 1$  if E-gene  $E_i$  shows in experiment  $k$  the same expression as in the negative controls.



# Likelihood

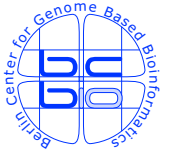


The silencing scheme  $\Phi$  allows **prediction** of E-gene states (when position is known).

We expect a number of **false positive and false negative** observations.

The **likelihood**  $P(D|\Phi, \Theta)$  is a product over atomic terms:

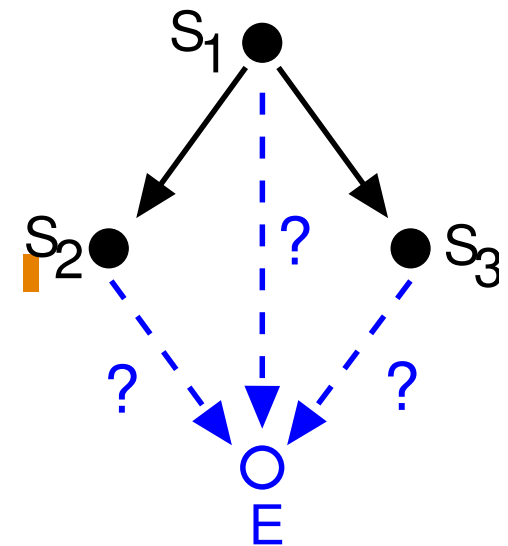
$$P(e_{ik}|\Phi, \theta_i = j) = \begin{cases} \frac{e_{ik} = 1}{\alpha} & \frac{e_{ik} = 0}{1 - \alpha} & \text{if } \Phi \text{ predicts } \mathbf{no\ effect} \\ \frac{e_{ik} = 1}{1 - \beta} & \frac{e_{ik} = 0}{\beta} & \text{if } \Phi \text{ predicts } \mathbf{effect} \end{cases}$$



# Marginal likelihood

Computation of likelihood requires that E-gene positions are known. In reality this is not true.■

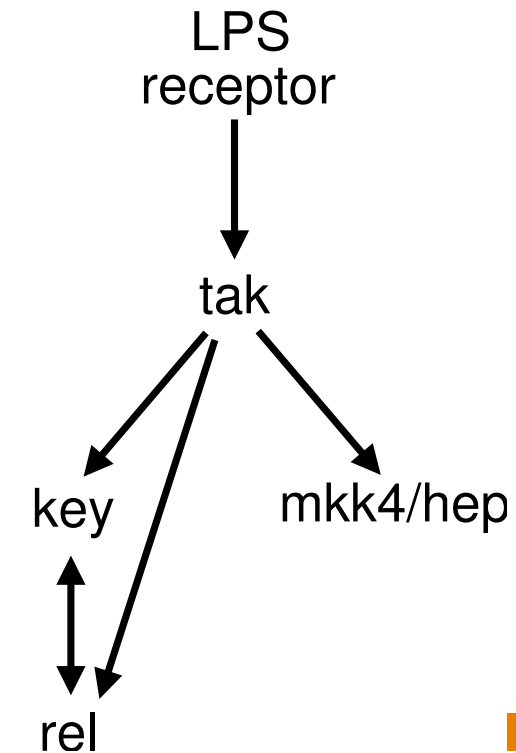
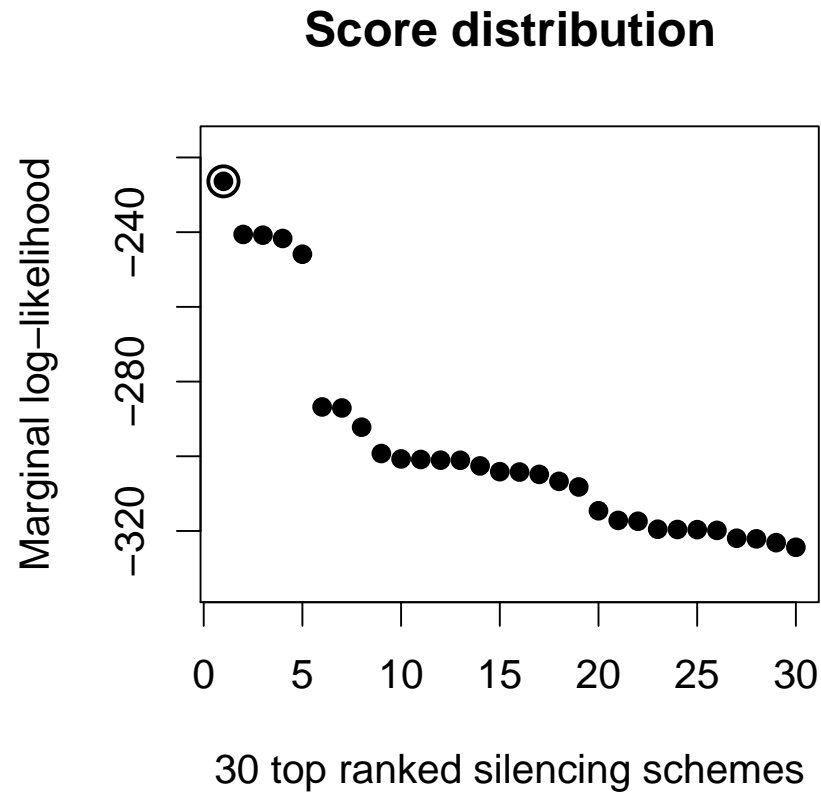
$$\begin{aligned} P(D|\Phi) &= \int P(D|\Phi, \Theta)P(\Theta|\Phi) d\Theta \blacksquare \\ &= \frac{1}{n^m} \prod_{i=1}^m \sum_{j=1}^n \prod_{k=1}^l P(e_{ik}|\Phi, \theta_i = j) \end{aligned}$$







# Application to *Drosophila* data

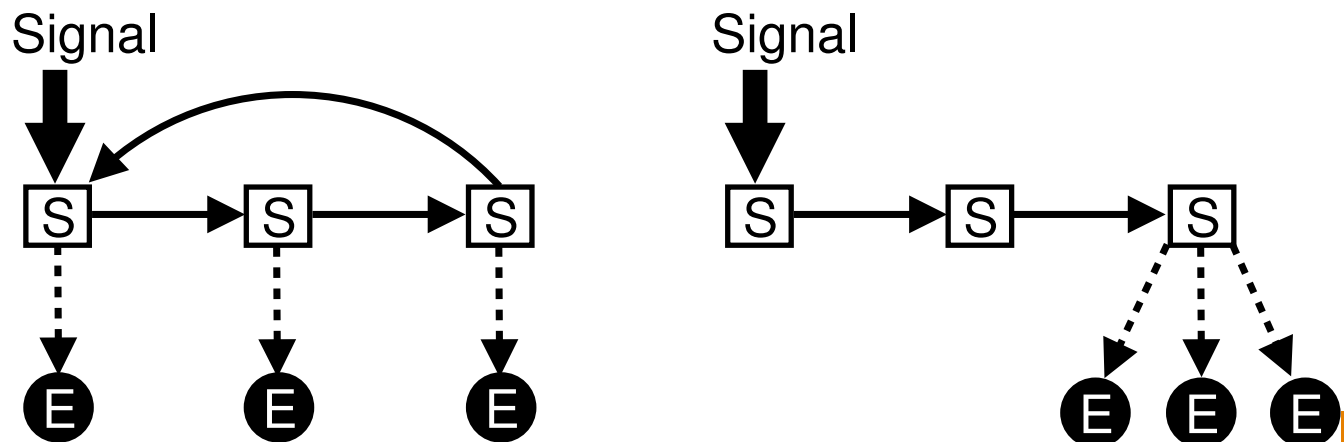




## Limits of identification

**Prediction equivalence** — Multiple pathway topologies result in the same silencing scheme, if they only differ in transitive edges.

**Likelihood equivalence** — Two hypotheses with different silencing schemes can produce identical data:





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## Conclusion

- The algorithm reconstructs pathway features from the nested structure of affected down-stream genes.■
- Pathway features are encoded as silencing schemes. They contain all information to predict a cell's behaviour to an external intervention. ■
- **Not shown:** in simulation studies we confirmed small sample size requirements and high accuracy.■
- Limitations only result from the information content of indirect observations.■



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## On true models

A quote from Edwards [3]:

*“Any method (or statistician) that takes a complex multivariate dataset and, from it, claims to identify one true model, is both naive and misleading.”*

What we have found is just

**a simple model consistent with the data.**



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# Graphical models in R

[www.r-project.org/gR](http://www.r-project.org/gR)

**ggm**: Gaussian Graphical Models

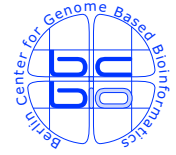
**deal**: Bayesian networks with mixed variables

[www.bioconductor.org](http://www.bioconductor.org)

**GeneTS**: large GGMs

[compdiag.molgen.mpg.de/software](http://compdiag.molgen.mpg.de/software)

**dcoex**: finding groups of differentially coexpressed genes



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# References

- [1] Katia Basso, Adam A Margolin, Gustavo Stolovitzky, *et al.* Reverse engineering of regulatory networks in human B cells. *Nat Genet*, Mar 2005.
- [2] Alberto de la Fuente, Nan Bing, Ina Hoeschele, and Pedro Mendes. Discovery of meaningful associations in genomic data using partial correlation coefficients. *Bioinformatics*, 20(18):3565–3574, 2004.
- [3] David Edwards. *Introduction to Graphical Modelling*. Springer, 2000.
- [4] MB Eisen, PT Spellman, PO Brown, and D Botstein. Cluster analysis and display of genome-wide expression patterns. *Proc Natl Acad Sci U S A*, 95(25):14863–8, Dec 1998.
- [5] Nir Friedman. Inferring Cellular Networks Using Probabilistic Graphical Models. *Science*, 303(5659):799–805, 2004.
- [6] Dennis Kostka and Rainer Spang. Finding disease specific alterations in the co-expression of genes. *Bioinformatics*, 20 Suppl 1:I194–I199, Aug 2004.
- [7] Paul M Magwene and Junhyong Kim. Estimating genomic coexpression networks using first-order conditional independence. *Genome Biol*, 5(12):R100, 2004.
- [8] Florian Markowetz, Jacques Bloch, and Rainer Spang. Non-transcriptional pathway features reconstructed from secondary effects of RNA interference. *Bioinformatics*, 2005.
- [9] Juliane Schäfer and Korbinian Strimmer. An empirical Bayes approach to inferring large-scale gene association networks. *Bioinformatics*, 21(6):754–64, Mar 2005.
- [10] PT Spellman, G Sherlock, MQ Zhang, *et al.* Comprehensive identification of cell cycle-regulated genes of the yeast *Saccharomyces cerevisiae* by microarray hybridization. *Mol Biol Cell*, 9(12):3273–97, Dec 1998.
- [11] Joshua M Stuart, Eran Segal, Daphne Koller, and Stuart K Kim. A gene-coexpression network for global discovery of conserved genetic modules. *Science*, 302(5643):249–55, Oct 2003.
- [12] Anja Wille and Peter Bühlmann. Tri-graph: a novel graphical model with application to genetic regulatory networks. Technical report, Seminar for Statistics, ETH Zrich, 2004.
- [13] Anja Wille, Philip Zimmermann, Eva Vranová, *et al.* Sparse graphical Gaussian modeling of the isoprenoid gene network in *Arabidopsis thaliana*. *Genome Biol*, 5(11):R92, 2004.